Estimating the magnetization distribution within rectangular rock samples

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Abstract Over the last decades, scanning magnetic microscopy techniques have been increasingly used in paleomagnetism and rock magnetism. Different from standard paleomagnetic magnetometers, scanning magnetic microscopes produce high-resolution maps of the vertical component of the magnetic induction field (flux density) on a plane located over the sample. These high-resolution magnetic maps can be used for estimating the magnetization distribution within a rock sample by inversion. Previous studies have estimated the magnetization distribution within rock samples by inverting the magnetic data measured on a single plane above the sample. Here we present a new spatial domain method for inverting the magnetic induction measured on four planes around the sample in order to retrieve its internal magnetization distribution. We have presumed that the internal magnetization distribution of the sample varies along one of its axes. Our method approximates the sample geometry by an interpretation model composed of a one-dimensional array of juxtaposed rectangular prisms with uniform magnetization. The Cartesian components of the magnetization vector within each rectangular prism are the parameters to be estimated by solving a linear inverse problem. Our method automatically deals with the averaging of the measured magnetic data due to the finite size of the magnetic sensor, preventing the application of a deconvolution before the inversion. Tests with synthetic data show the performance of our method in retrieving complex magnetization distributions even in the presence of magnetization heterogeneities. Moreover, they show the advantage of inverting the magnetic data on four planes around the sample and how this new acquisition scheme improves the estimated magnetization distribution within the rock sample. We have also applied our method to invert experimentally measured magnetic data produced by a highly magnetized synthetic sample that was manufactured in the laboratory. The results show that even in the presence of apparent position noise, our method was able to retrieve the magnetization distribution consistent with the isothermal remanent magnetization induced in the sample.

1. Introduction

Based on the estimation of the total (bulk) remanent magnetization (RM), the standard paleomagnetic techniques have been developed for more than 40 years in studies of rock magnetism, magnetostratigraphy, and paleogeographic reconstructions [e.g., Van Der Voo, 1993]. The evolution of the paleomagnetic techniques is directly linked to the development of rock magnetometers and data processing methods.

Since its beginning in the 1950s, paleomagnetism has seen at least three generation of rock magnetometers [Turner et al., 2015]: astatic systems, spinner magnetometers, and superconducting magnetometers. Astatic system was used during the earliest paleomagnetic studies, such as Blackett [1952] and Pozzi and Thellier [1963], and it was quickly replaced by spinner magnetometers [Nagata, 1961]. Spinner magnetometers are most used devices in the paleomagnetic laboratories but they have been systematically replaced by superconducting magnetometers. Superconducting rock magnetometers have been developed since 1970s [Dodson et al., 1974; Goree and Fuller, 1976], reaching notable sensitivity and quickness of measurements on the last two decades [Kirschvink et al., 2015]. This kind of rock magnetometer can measure both discrete samples (single sample mode) and long-core samples (pass-through mode), with a good spatial accuracy (~1 cm). The long-core samples or u-channel samples are essential to investigations of magnetostratigraphy and environmental magnetism, such as those ones obtained at IODP (International Ocean Drilling Project).
In terms of data processing, both spinner and SQUID (discrete samples) are similar. For pass-through mode, however, the data processing is little bit different; the measurements are naturally smoothed and have a direct dependence with the sensor response. To overcome this constraint, some studies have used deconvolution algorithms to recover the paleomagnetic signal [Jackson et al., 2010; Oda and Xuan, 2014]. Even though the deconvolution methods have been improved on the last years [Jackson et al., 2010; Lascu et al., 2012], Lascu et al. [2012] have noticed that there is a resolution constraint related with sensor response. They have compared the measurements of long-core with the discrete (slices samples) synthetic samples and they realize that the sensor setting imposes a spatial resolution limit of 1 cm. Along these lines this method has constraint in studies that need very high spatial or temporal resolution, such as the secular variation recorded in stalagmites or manganese crusts.

In this context, new methods have been developed over the last decades to get around these limitations, in special, the Scanning Magnetic Microscopy (SMM) techniques [Kirschvink et al., 2008; Oda et al., 2011; Fu et al., 2014]. Different from standard paleomagnetic magnetometers, the SMM devices produce high-resolution maps of the vertical component of the magnetic induction field on a planar surface located over the sample. Usually, the scanning magnetic microscopes based on superconducting quantum interference device (SQUID) sensors achieve the best field sensitivity for geoscientific research. However, these devices are very expensive to build/operate and require specific technologies to keep the sensor at cryogenic temperatures and also present a limited sensor-to-sample distance \( \approx 100 \, \mu m \) [Baudenbacher et al., 2003; Fong et al., 2005]. Despite these technical difficulties, SQUID sensors have been successfully applied in several paleomagnetic studies [Gattacceca et al., 2006; Weiss et al., 2007a; Oda et al., 2011; Fu et al., 2014; Kirschvink et al., 2015]. To avoid the problems related with the SQUID complexity, numerous studies have been attempted to develop alternative low-cost high-performance SMM instruments based on magneto-impedance (MI) [Uehara and Nakamura, 2007, 2008], giant magnetoresistance (GMR) [Hankard et al., 2009], and magnetic tunnel junction (MTJ) [Lima et al., 2014] sensors. Compared with SQUIDs, these alternative sensors are less sensitive. On the other hand, they are generally easier to operate, operated at room temperature, and may achieve higher spatial resolution and also a reduction in the sensor-to-sample distance.

Paleomagnetic techniques usually separate the original rock sample mechanically and estimate the magnetization of resultant small subsamples. In contrast, SMM allows a nondestructive characterization of the internal magnetization distribution of rock samples by inverting high-resolution magnetic data that is measured around the sample. SMM may provide huge data sets and, consequently, the inverse problems aiming at estimating the magnetization within the samples require efficient matrix algorithms. It is also well known that inverse problems aiming at estimating the magnetization distribution in planar or 3D, e.g., cylindrical and rectangular, rock samples are generally nonunique owing to an infinite number of magnetization distributions that produce the same observed field. Moreover, the observed magnetic data are always noise-corrupted and is measured by magnetic sensors having limited sensitivity [Egli and Heller, 2000; Baratchart et al., 2013; Lima et al., 2013], contributing to the ill-conditioning of such inverse problems. The inherent ill-conditioning of such inverse problems can be narrowed, for example, by introducing a priori information regarding the magnetization distribution and/or by optimizing the geometry of the data acquisition. The introduction of a priori information aiming at constraining the possible estimated magnetization distributions and also making them stable to small changes in the observed data is generally called regularization [Tikhonov and Arsenin, 1977; Engl et al., 1996; Aster et al., 2005]. The regularization in the wavenumber domain is generally more tricky than in the space domain. On the other hand, methods in space domain require the solution of large-scale linear systems whereas those ones in wavenumber domain take advantage of the Fast Fourier Transform (FFT) algorithms.

Generally, SMM paleomagnetic studies estimate the approximately 2D magnetization distribution within planar rock samples (usually thin sections) by inverting the magnetic data measured on a parallel plane located over it. The pioneer study of Egli and Heller [2000] presents a method, in the wavenumber domain, aiming at retrieving the particular magnetization component that is perpendicular to the planar sample. These authors adapted the method proposed by Mareschal [1985] and formulated their problem as a two-dimensional deconvolution, which is solved by applying the FFT. By following a similar approach in the wavenumber domain, Lima et al. [2013] presented an efficient method attempted to estimate the magnetization intensity distribution within a planar rock sample having a previously
defined constant magnetization direction. The method developed by Lima et al. [2013] also uses two-dimensional signal processing methods to regularize the inverse problem, tame noise amplification, and improve nonnegativity.

A different approach proposed by Weiss et al. [2007b] in space domain approximates the sample by a discrete set of dipoles. The Cartesian components of the dipoles approximating the sample are estimated by iteratively solving a large linear inverse problem. This method is an adaptation of the well-known equivalent layer technique, which has long been applied to processing potential-field data in space domain [Dampney, 1969; Emilia, 1973; Von Frese et al., 1981; Hansen and Miyazaki, 1984; Silva, 1986; Leão and Silva, 1989; Cordell, 1992; Mendonça, 1992; Mendonça and Silva, 1994, 1995; Guspi and Novara, 2009; Li and Oldenburg, 2010; Barnes and Lumley, 2011; Oliveira Jr. et al., 2013; Kara et al., 2014; Li et al., 2014; Barnes, 2014]. The study developed by Weiss et al. [2007b] represents the first space-domain technique for inverting SMM data. Usui et al. [2012] presented a hybrid method combining useful features of space- and wavenumber-domain techniques. They described the magnetization distribution in a planar rock sample as a weighted average according to the Backus-Gilbert formulation. In order to overcome the computational cost of the Backus-Gilbert method, Usui et al. [2012] implemented the subtractive optimally localized averages (SOLA) method, which was originally proposed to solve large-scale inverse problems in helioseismic. The SOLA method approximates some matrix computations by using the FFT, but does not transform the magnetic data to the Fourier domain.

Here we propose a new spatial domain inversion method to invert the magnetic induction measured on four orthogonal planes around the sample in order to retrieve its internal magnetization distribution. Our method approximates the sample by an interpretation model composed of a one-dimensional array of juxtaposed rectangular prisms with uniform magnetization. The number of rectangular prisms making up the interpretation model is specified by the interpreter and the Cartesian components of the magnetization vector within each rectangular prism are the parameters to be estimated by solving a linear inverse problem.

By imposing this finite one-dimensional magnetization distribution, we constraint the mathematically possible solutions, which results in a regularization by discretization [Engl et al., 1996; Aster et al., 2005]. Moreover, the use of magnetic data measured on more than one plane around the sample adds independent information about its internal magnetization distribution, which also contributes to stabilize the inverse problem. We also presents a new approach for automatically including the averaging of the magnetic data over the finite active area of the magnetic sensor. This new approach precludes the application of a deconvolution for removing this averaging effect before the inversion. In comparison with previous methods, the method proposed here does not deal with large-scale inverse problems because our interpretation model is described by a relatively few number of parameters.

We have inverted magnetic data produced by numerical simulations with the purpose of illustrating not only the good performance of our method in ideal cases, but also how the estimated magnetization distribution obtained by our method can be negatively impacted by the presence of position noise, errors in the sensor-to-sample distance, and preprocessing errors in the observed data. Our results also show the benefit of inverting a full data set formed by magnetic data measured on four planes around the sample. Tests with synthetic data simulating a real ferro-manganese crust with magnetization heterogeneities show that our method can be used, for example, in fine-scale magnetostratigraphic studies. We have also shown the results obtained by applying our method to invert experimentally measured magnetic data produced by a highly magnetized synthetic sample that was manufactured in laboratory. The results show that our method is able to estimate a magnetization distribution that is consistent with the Isothermal Remanence Magnetization (IRM) induced in the sample, even by inverting a magnetic data set that is contaminated by apparent position noise and were measured by a custom-made magnetometer based on a Hall sensor.

2. Methodology

2.1. Observed Data Vector

Let us consider a rectangular rock sample with side lengths equal to \( L_x \), \( L_y \), and \( L_z \) along, respectively, the \( x \), \( y \) and \( z \)-axes of a Cartesian coordinate system whose origin coincides with the center of the sample (Figure 1). This coordinate system is conveniently called “main coordinate system” (MCS). We assume that the
internal magnetization distribution of the sample varies along the x axis of the MCS. We also considered four mutually orthogonal planes that are located at the same distance \( h \) from the sample surface and are identified by an index \( \alpha = 0, 1, 2, 3 \) (Figures 2 and 3). On each plane, there are \( N \) measurements of a specific \( \beta \) component of the magnetic induction, \( \beta = yz \), which is perpendicular to the sample surface and is referred to the MCS.

Let \( \mathbf{d}_\alpha^\beta \), \( \beta = y, z, \alpha = 0, 1, 2, 3 \), be \( N \times 1 \) vectors whose \( i \)th element is the \( \beta \) component of the magnetic induction which is measured at the observation point \( (x_i, y_i, z_i) \), \( i = 1, \ldots, N \), on the plane \( \alpha \). For convenience, these vectors are all grouped into the \( 4N \times 1 \) observed data vector \( \mathbf{d} \) given below:

\[
\mathbf{d} = \begin{bmatrix} \mathbf{d}_0^y \\ \mathbf{d}_1^y \\ \mathbf{d}_2^y \\ \mathbf{d}_3^y \\ \mathbf{d}_0^z \\ \mathbf{d}_1^z \\ \mathbf{d}_2^z \\ \mathbf{d}_3^z \end{bmatrix}
\] (1)

In practical situations, however, we are not able to measure the magnetic induction produced by the sample at a point. Due to the finite size of the magnetic sensors, what is actually measured is the magnetic induction averaged on the area (or volume) of the magnetic sensor. In the section 2.3, we present an analysis of this averaging effect.

2.2. Transformations From the Local Coordinate Systems (LCs) to the MCS

The Cartesian coordinates as well as the magnetic induction components in the observed data vector \( \mathbf{d} \) (equation (1)) are referred to the MCS (Figure 1). However, the measurements are taken in a different local coordinate system (LCS) for each observation plane (Figures 3b–3e) and must be subsequently converted to the MCS (Figure 3a). This coordinate transformation is needed because the usual magnetic sensors sense the vertical component of the magnetic induction produced by the sample on a plane located over it. For this reason, the magnetic data on the four observation planes (Figure 2) are obtained by successively rotating the sample through \( 90^\circ \) intervals around its major axis (x axis in Figure 1). By repeating this rotating procedure and maintaining the same distance \( h \) between the observation plane and the sample surface, it
is possible to obtain the magnetic induction on the four observation planes (Figure 2). The LCS on each plane has axes $x', y'$, and $z'$ (Figures 3b–3e), where the $x'$ axis coincides with the $x$ axis of the MCS.

Note that this rotating procedure provides, on each observation plane, the component of the magnetic induction along the $z'$ axis of the respective LCS (Figures 3b–3e). For convenience, let us denote the measured $z'$ component data as a $4N \times 1$ vector given by

$$d' = \begin{bmatrix} d'_{z}^{0} \\ d'_{z}^{1} \\ d'_{z}^{2} \\ d'_{z}^{3} \end{bmatrix},$$

where $d'_{z}^{x}$, $x=0, 1, 2, 3$, are $N \times 1$ vectors whose $i$th element is the $z'$ component of the magnetic induction which is measured at the $i$th observation point, $i=1, \ldots, N$, on the plane $x$ (Figure 3). It is easier, however, converting the measured data $d'$ (equation (2)), referred to the LCSs, to the observed data vector $d$ (equation (1)), which is referred to the MCS.

The geometrical relationships between the LCSs and MCS are shown in Figure 3. Table 1 shows how to perform this coordinate transformation. Let us consider, for example, the measurements obtained on the observation plane $x = 1$ (Figures 2b and 3c). According to this table, the opposite of the measured $z'$ component data correspond to the $y$ component of the magnetic induction that would be measured in the MCS; the opposite of the Cartesian coordinates along the $z'$ axis correspond to the Cartesian coordinates along the $y$ axis in the MCS; the Cartesian coordinates along the $y'$ axis correspond to the Cartesian coordinates along the $z$ axis in the MCS. The other lines of Table 1 contain the relationships used to convert the data obtained on the other planes. Hereafter, it is implicit that all quantities with prime ($) are referred to the LCSs while all the quantities without prime ($) are referred to the MCS.

2.3. Sensor Active Area
As pointed out at the end of the section 2.1, magnetic sensors do not measure the magnetic induction field at a point. They measure instead the field averaged over the sensor active area (or volume), which is equivalent to a convolution of the magnetic induction field with the sensor answer function [Roth et al., 1989; Egli and Heller, 2000; Lima et al., 2013]. The sensor-to-sample distance and the effective area (or volume) are important limiting factors determining the spatial resolution of magnetic sensors.

The magnetic sensors used in paleomagnetic applications have different active area (or volume). Uehara and Nakamura [2007, 2008] presented a magneto-impedance (MI) sensor whose cylindrical active volume has 36 $\mu$m in diameter and 5500 $\mu$m in length, which limits its spatial resolution if compared with the giant magnetoresistance (GMR) sensor presented by Hankard et al. [2009] and the magnetic tunnel junction (MTJ) sensor presented by Lima et al. [2014], for example. These two sensors are considered planar and have active areas of, respectively, $9 \mu$m $\times$ $36 \mu$m and $\approx 4 \mu$m $\times$ $2 \mu$m. Finally, Fong et al. [2005] used a bare SQUID design with an active diameter of 120 $\mu$m. In this case, the sensor is also considered planar.

In the following, we present an analysis of the averaging effect produced by the finite size of sensors. For convenience, we limit our study to square-planar sensors which are parallel to the sample surface and have
an active area $A$ (Figure 5a). Let us first assume that the sensor detects, at the $i$th observation point $(x_i', y_i', z_i')$, $i=1, \ldots, N$, located on an observation plane $x = x_0$, the $z'$ component of the magnetic induction (flux density) (Figure 3) that is threaded through the active area $A$ (light gray square in Figure 5a) of the sensor. For convenience, let us denote this observed $z'$ component as $d'_i = d'(x_i', y_i', z_i')$, which represents the $i$th element of a particular vector $\mathbf{d}_z'$ (equation (2)), depending on the observation plane $x$ on which the measurement is done. This observed data $d'_i$ can then be described as a two-dimensional convolution given by

$$d'_i = \frac{1}{A} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} w(\eta, \zeta) d'(x'_0 - \eta, y'_0 - \zeta, z'_0) \, d\eta d\zeta,$$

where $w(\eta, \zeta)$ is the sensor answer function [Egli and Heller, 2000] and $d'(x'_0 - \eta, y'_0 - \zeta, z'_0)$ represents the theoretical values of the $z'$ component of the magnetic induction produced by the rock sample at the points located on the observation plane $x$ with constant $z'_0$ (Figure 3). Here we do not consider the dependence with the sensor-to-sample distance $h$ (Figure 5a). In this case, the sensor answer function is given by [Egli and Heller, 2000]

$$w(x'_i, y'_i) = \begin{cases} 1/A, & (x'_i, y'_i) \in A \\ 0, & (x'_i, y'_i) \notin A \end{cases}.$$  

By substituting this sensor answer function into equation (3), we obtain

$$d'_i = \frac{1}{A} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} d'(x'_0 - \eta, y'_0 - \zeta, z'_0) \, d\eta d\zeta.$$  

Equation (5) has an analytical solution in the wavenumber domain [Egli and Heller, 2000]. However, we opted for using an approximated solution in the space domain. This approximation consists in discretizing the integral 5 by using a midpoint rule on a regular grid of $Q_x \times Q_y$ points $(x'_j, y'_j) \in A$, $j=1, \ldots, Q$, where $Q = Q_x Q_y$ (open circles in Figure 5b). According to this discretization, the active sensor area $A$ is divided into a regular grid of square cells with area $\Delta A = A/Q$ (Figure 5b). The approximated integral is given by

$$d'_i \approx \frac{1}{A} \sum_{j=1}^{Q} d'_j \Delta A,$$

where $d'_j$ is the theoretical value of the $z'$ component of the magnetic induction produced by the rock sample at the point $(x'_j, y'_j)$, $j=1, \ldots, Q$ (open circles in Figure 5b).

Equation (6) shows that the $z'$ component $d'_i$ measured at the $i$th observation point on a plane $x$ is approximately given by the averaged magnetic induction produced by the rock sample over the active area $A$ of the sensor. Consequently, the elements of the observed data vector $\mathbf{d}$ (equation (1)) also represent averaged components of the magnetic induction over the active area $A$ of the magnetic sensor.

### 2.4. Preprocessing

Our method presumes that the edges of the sample are aligned with the axes of all LCSs (Figures 3b–3e) and also that the center of the sample must be placed right below the origin of all LCSs. Notwithstanding, these conditions are not necessarily satisfied in practical situations. Figure 4a illustrates a situation in which these conditions are satisfied. On the other hand, Figure 4b shows a situation in which neither the horizontal coordinates of the center of the sample (black dot) coincide with the origin of the LCS (open dot), nor the edges of the sample are aligned with the axes $x'$ and $y'$ of the LCS. In this case, it is necessary to correct the coordinates $x'$ and $y'$ of the magnetic data on the observation plane with respect to the sample in order to positioning the sample according to the Figure 4a. This correction must be applied to the magnetic data obtained on the four observation planes.
On each observation plane, let us first denote by \( x'_0 \) and \( y'_0 \) the horizontal coordinates of the center of the observation plane (represented by the open dot in Figure 4b). Then, the horizontal coordinates \( \tilde{x}'_0 \) and \( \tilde{y}'_0 \) of each magnetic data on the observation plane are corrected by applying the following transformation:

\[
\begin{bmatrix}
\tilde{x}'_0 \\
\tilde{y}'_0 
\end{bmatrix} = R \begin{bmatrix}
x' - x'_0 \\
y' - y'_0 
\end{bmatrix} + \begin{bmatrix}
\Delta x' \\
\Delta y' 
\end{bmatrix},
\]

(7)

where \( \Delta x' \) and \( \Delta y' \) are the horizontal displacements of the center of the sample with respect to the origin of the respective LCS (Figure 4b),

\[
R = \begin{bmatrix}
\cos \theta & \sin \theta \\
-\sin \theta & \cos \theta 
\end{bmatrix},
\]

(8)

and \( \theta \) (Figure 4b) is the angle between the edges of the sample and the horizontal axes \( x' \) and \( y' \) of the respective LCS. After applying this transformation for correcting the horizontal coordinates \( \tilde{x}'_0 \) and \( \tilde{y}'_0 \) of each magnetic data on an observation plane, we obtain the corrected coordinates \( x'_0 \) and \( y'_0 \), which are placed on a different area represented by the gray rectangle in Figure 4c.

This transformation must be applied to the magnetic data on the four observation planes, before the previously described transformations from the LCSs to the MCS. We would like to stress that the transformation described in equations (7) and (8) requires the knowledge of the angle \( \theta \) and also the displacements \( \Delta x' \) and \( \Delta y' \) (Figure 4b). In practical situations, these values are easily estimated by trial-and-error. Finally, on each observation plane, the magnetic data are subtracted from its mean value in order to remove the effect of a possible (weak) interfering field.

2.5. Interpretation Model and Forward Problem

We consider that the rock sample can be approximated by an interpretation model consisting of \( P \) uniformly magnetized prisms, which are juxtaposed along the \( x \) axis (Figure 1). The \( k \)th prism, \( k = 1, \ldots, P \), has the same side lengths \( L_y \) and \( L_z \) of the sample along, respectively, the \( y \) and \( z \) axes (Figure 1). However, its side length along the \( x \) axis is specified by the interpreter, so that the total side length of the interpretation model along this axis is equal to \( L_x \) (Figure 1).
Let us define the $3P \times 1$ parameter vector $m$ as follows:

$$m = \begin{bmatrix} m_1 \\ \vdots \\ m_P \end{bmatrix},$$

(9)

where

$$m^k = \begin{bmatrix} m^k_x \\ m^k_y \\ m^k_z \end{bmatrix}$$

(10)

is a $3 \times 1$ vector containing the Cartesian components $m^k_x$, $m^k_y$, and $m^k_z$ (in A/m) of the magnetization vector of the $k$-prism, $k=1,\ldots,P$.

Let us also define $N \times 1$ coordinate vectors $x^i$, $y^i$, and $z^i$, whose $ith$ elements are the coordinates $x_i^i$, $y_i^i$, and $z_i^i$, respectively, of the $ith$ observation point on the plane $\pi$ (black circle in Figure 5b). Additionally, consider the $QN \times 1$ coordinate vectors $x^i$, $y^i$, and $z^i$. The first $Q$ elements of these vectors are, respectively, the coordinates $x_i^i$, $y_i^i$, and $z_i^i$ of a regular grid of points (open circles in Figure 5b) centered at the observation point $(x_i^i, y_i^i, z_i^i)$ (black circle in Figure 5b). Similarly, the next set of $Q$ elements of these vectors are the coordinates of a regular grid of $Q$ points centered at the observation point $(x_{i+Q}^i, y_{i+Q}^i, z_{i+Q}^i)$ and so on. The Cartesian coordinates forming the vectors $x^i$, $y^i$, and $z^i$ and $x_i^i$, $y_i^i$, and $z_i^i$ are referred to the MSC (Figure 3).

Once the coordinate vectors were defined, we define the $\beta$ component, $\beta = y, z$, of the magnetic induction produced by the $k$-prism (in $\mu T$), at the observation points $x^i$, $y^i$, and $z^i$, are grouped into $N \times 1$ vectors given by

$$b_{\beta}(x^i, y^i, z^i, m^k) = AM_{\beta}(x^i, y^i, z^i)m^k,$$

(11)

where $A$ is a $N \times QN$ sparse matrix whose element $ij, i=1,\ldots,N, j=1,\ldots,QN$, is given by

$$[A]_{ij} = \begin{cases} 1/Q, & (i-1)Q < j \leq iQ \\ 0, & \text{otherwise} \end{cases}$$

(12)

and

$$M_{\beta}(x^i, y^i, z^i) = \begin{bmatrix} \frac{\partial}{\partial x} \phi^k \frac{\partial}{\partial x} \phi^k \\ \frac{\partial}{\partial y} \phi^k \\ \frac{\partial}{\partial z} \phi^k \frac{\partial}{\partial x} \phi^k \\ \vdots \\ \frac{\partial}{\partial y} \phi^k \\ \frac{\partial}{\partial z} \phi^k \frac{\partial}{\partial y} \phi^k \\ \frac{\partial}{\partial z} \phi^k \end{bmatrix}$$

(13)

is a $QN \times 3$ matrix whose elements $\frac{\partial}{\partial x} \phi^k$, $\frac{\partial}{\partial y} \phi^k$, and $\frac{\partial}{\partial z} \phi^k, j=1,\ldots,QN$, are second derivatives, evaluated at the points $x^i$, $y^i$, and $z^i$, of the function

$$\phi^k(x, y, z) = C_m \iiint_{\mathcal{Q}} \frac{d\epsilon \, d\zeta \, d\eta}{\sqrt{(x-e)^2 + (y-\zeta)^2 + (z-\eta)^2}}$$

(14)

with respect to the variables $x, y$, and $z$. In equation (14), $C_m = 10^3 \mu_0 / 4\pi$, $\mu_0$ = $4 \pi 10^{-7}$ H/m is the magnetic permeability of the free space and the integral is evaluated over the volume $\phi^k$ of the $k$th prism. In this work, the second derivatives in the matrices $M_{\beta}(x^i, y^i, z^i)$ (equation (13)) are calculated by using the formulas presented by Nagy et al. [2000]. Notice that the vectors $b_{\beta}(x^i, y^i, z^i, m^k)$ (equation (11)) contain averaged values of the $\beta$ components of the magnetic induction produced by the interpretation model at
the observation points located at the four planes $\alpha=0, 1, 2, 3$. The averaging is imposed by the matrix $A$ defined in equation (12).

From equation (11), we define the $N \times 1$ vectors $B^m_{\beta}(m)$ containing the $\beta$ component of the magnetic induction produced by all prisms forming the interpretation model on each plane $\alpha$ as follows:

$$B^m_{\beta}(m) \equiv \sum_{k=1}^{P} b_{\beta}(x^k, y^k, z^k, m^k).$$  (15)

By grouping all these vectors $B^m_{\beta}(m)$, we define the $4N \times 1$ theoretical-predicted data vector

$$B(m) = \begin{bmatrix} B^m_{y1}(m) \\ B^m_{y1}(m) \\ B^m_{z1}(m) \\ B^m_{z1}(m) \end{bmatrix}. $$  (16)

Finally, by substituting equation (11) into equation (15) and rearranging the terms within the summation, we can conveniently rewrite the predicted data vector $B(m)$ (equation (16)) as follows:

$$B(m) = Mm. $$  (17)

where $M$ is a $4N \times 3P$ partitioned matrix given by

$$M = \begin{bmatrix} AM^y_0(x^0, y^0, z^0) & \cdots & AM^y_0(x^P, y^0, z^0) \\ AM^y_0(x^0, y^1, z^1) & \cdots & AM^y_0(x^P, y^1, z^1) \\ AM^z_0(x^2, y^2, z^2) & \cdots & AM^z_0(x^2, y^2, z^2) \\ AM^z_0(x^3, y^3, z^3) & \cdots & AM^z_0(x^3, y^3, z^3) \end{bmatrix}. $$  (18)

where $A$ is defined in equation (12) and $M_0^\beta(x^k, y^k, z^k), \beta = y, z, \alpha = 0, 1, 2, 3, k = 1, \ldots, P$ in equation (13).

Notice that the forward problem presented in this subsection can be easily modified to the case in which the magnetic data are measured on less than four observation planes. Besides, this forward problem describes the averaged magnetic induction produced by the interpretation model at the observation points. The averaging is calculated over a regular grid of $Q$ points on the active area $A$ of the magnetic sensor (Figure 5).

### 2.6. Inverse Problem

By presuming that the rectangular rock sample can be approximated by our previously described interpretation model, we define the linear inverse problem of estimating its internal magnetization distribution as a constrained optimization problem. This optimization problem consists in estimating a specific parameter vector $m = m^*$ (equation (9)) that minimizes the goal function

$$\Gamma(m) = ||d - B(m)||^2 + \mu_0 f_0 ||Mm||^2, $$  (19)

where $|| \cdot ||_2$ stands for Euclidian (or $\ell_2$) norm, $d$ and $B(m)$ are, respectively, the observed and predicted data vectors (equations (1) and (17)), $f_0$ is defined as the ratio

$$f_0 = \frac{tr(M^T M)}{3P}, $$  (20)

where $tr(\cdot)$ stands for the trace, $\mu_0$ is a positive number called regularization parameter, and $R$ is a $3P \times 3 \times 3$ sparse matrix whose $ij$th-element is given by

$$[R]_{ij} = \begin{cases} 1, & i=j \\ -1, & j = i+3 \\ 0, & \text{otherwise} \end{cases}. $$  (21)

In equation (19), $f_0$ is used with the purpose of constraining the practical range of $\mu_0$, which controls the trade-off between the first and second terms at the right side. The first term represents the data fit and the
second term is the well-known first-order Tikhonov regularization function [Aster et al., 2005], which imposes “smoothness” to the estimated parameters. The least-squares estimate of the parameter vector \( \mathbf{m} = \mathbf{m}^\dagger \) minimizing the goal function (equation (19)) is given by

\[
\mathbf{m}^\dagger = (\mathbf{M}^\top \mathbf{M} + \mu_0 \mathbf{R}^\top \mathbf{R})^{-1} \mathbf{M}^\top \mathbf{d}.
\]

We consider that the estimated parameter vector \( \mathbf{m}^\dagger \) given by equation (22) approximates the real magnetization distribution within the rectangular rock sample.

### 2.7. Software Implementation


The computational experiments (e.g., data processing, synthetic tests, and real data application) were performed in Jupyter (formerly IPython) notebooks [Pérez and Granger, 2007, http://jupyter.org]. The notebook files combine the source code, the results and figures generated by the code and rich text to explain the analysis.

All source code, Jupyter notebooks, data, and results can be found at the online repository https://github.com/pinga-lab/magnetization-rock-sample.

### 3. Numerical Simulations

#### 3.1. Highly Magnetized Sample

We have applied our method to invert the synthetic data produced by a simulated rock sample (not shown) with \( L_x = 16 \text{ mm} \) and \( L_y = L_z = 3 \text{ mm} \), according to the Figure 1. This synthetic sample is formed by \( P = 16 \) juxtaposed prisms along the \( x \) axis, where each prism has a uniform and high magnetization (red dots in Figures 7, 8 and 11). In all tests presented in this section, the simulated magnetic sensor has a square active area \( A = (300 \mu \text{m})^2 \), which was discretized into a regular grid of \( 7 \times 7 \) points, totaling \( Q = 49 \) (Figure 5). The magnetic data produced by the synthetic sample on each plane \( z = 0, 1, 2, 3 \) were calculated at \( N = 4 \), 284 points, on a set of 42 regularly spaced lines that are parallel to the \( x \) axis, each one with 102 regularly spaced points. The data were also contaminated with a pseudo-random Gaussian noise having null mean and standard deviation equal to 30 \( \mu \text{T} \). These data simulate an observed data set. Besides, we have simulated the misalignment problems described in the section 2.4. Table 2 shows the parameters \( \theta, \Delta \chi, \) and \( \Delta \lambda^\dagger \) (Figure 4b) representing these misalignments in the magnetic data produced by the synthetic sample on the planes \( z = 0, 1, 2, 3 \).

#### 3.1.1. Validation Test

In this test, the simulated noise-corrupted data were calculated by keeping the distance \( h \) (Figure 3) between all the planes and the surface of the sample equal to 500 \( \mu \text{m} \). The noise-corrupted magnetic data were properly corrected from the misalignment problems by using the parameters shown in Table 2 and the equations (7) and (8). These magnetic data (Figures 6a, 6d, 6g, and 6j) were inverted by our method in order to retrieve the internal magnetization distribution of the synthetic sample (red dots in Figures 7 and 8), and 11). To do that, we used an interpretation model which has side lengths \( L_x, L_y, \) and \( L_z \) equal to the true ones and is formed by the same number of prisms (\( P = 16 \)) as the synthetic sample.

Figure 7 shows the four estimated magnetization distributions obtained by our method. The one obtained by inverting the magnetic data on the four planes \( z = 0, 1, 2, 3 \) is represented by the black dots and dashed lines. The three sets of blue dots and dashed lines represent the estimated magnetization distribution obtained by inverting the data located on the plane \( z = 0 \), planes \( z = 0, 1, \) and planes \( z = 0, 1, 2 \), which are plotted, respectively, with high, intermediate, and low transparency. All these estimates were obtained by using a regularization parameter \( \mu_0 = 1 \times 10^{-10} \) (equation (19)). We can see that the estimated magnetization distribution obtained by inverting the full data set (black dots and black dashed lines in Figure 7) is the best one in retrieving the true magnetization distribution (red dots and red lines in Figure 7). This estimated magnetization distribution yields a predicted data (Figures 6b, 6e, 6h, and 6k) that is very close to the observed data (Figures 6a, 6d, 6g, and 6j). The normalized histograms of the residuals between the
Validation test. (a, d, g, j) Noise-corrupted magnetic data produced by the synthetic sample (not shown) on the observation planes \( z = 0, 1, 2, \) and 3, respectively. (b, e, h, k) Predicted data produced by the estimated magnetization distribution on the observation planes \( z = 0, 1, 2, \) and 3, respectively. The estimated magnetization distribution was obtained by inverting the magnetic data located on the four observation planes. (c, f, i, l) Normalized histograms of the residuals between the predicted data shown in Figures 6b, 6e, 6h, and 6k, and the noise-corrupted magnetic data shown in Figures 6a, 6d, 6g, and 6j. The normalization consists in subtracting from the residuals its sample mean \( \mu \) and dividing the result by its sample standard deviation \( \sigma \). The values are in \( \mu T \).

Figure 6. Validation test. (a, d, g, j) Noise-corrupted magnetic data produced by the synthetic sample (not shown) on the observation planes \( z = 0, 1, 2, \) and 3, respectively. (b, e, h, k) Predicted data produced by the estimated magnetization distribution on the observation planes \( z = 0, 1, 2, \) and 3, respectively. The estimated magnetization distribution was obtained by inverting the magnetic data located on the four observation planes. (c, f, i, l) Normalized histograms of the residuals between the predicted data shown in Figures 6b, 6e, 6h, and 6k, and the noise-corrupted magnetic data shown in Figures 6a, 6d, 6g, and 6j. The normalization consists in subtracting from the residuals its sample mean \( \mu \) and dividing the result by its sample standard deviation \( \sigma \). The values are in \( \mu T \).
predicted and observed data (Figures 6c, 6f, 6i, and 6l) show sample means $\mu$ and sample standard deviations $\sigma$ very close to that of the pseudo-random noise contaminating the simulated magnetic data.

These results show the advantage of inverting the full data set (data on the four planes) and also the good performance obtained by our method if all the premisses about the magnetic data and the sample are not violated.

### 3.1.2. Preprocessing Errors

As in the previous test, the simulated noise-corrupted data used here were calculated by keeping the distance $h$ (Figure 3) between all the planes and the surface of the sample equal to 500 $\mu$m. On the other hand, we have not corrected the misalignment problems by using the parameters shown in Table 2 and the equations (7) and (8). We have also introduced a different constant bias into the misaligned magnetic data on each plane with the purpose of simulating errors in the preprocessing stage. These biases are equal to $-26,000$, $-10,000$, $-26,000$, and $-23,000$ $\mu$T for the magnetic data on the planes $x = 0$, 1, 2, and 3, respectively. As in the previous test, we applied our method by using an interpretation model with side lengths $L_x$, $L_y$, and $L_z$ equal to the true ones and with the same number of prisms ($P = 16$) as the synthetic sample along the $x$ axis.

Figure 8 confirms the superiority of the estimated magnetization distribution obtained by inverting the full data set (black dots and black dashed lines) over the remaining three (blue dots and blue dashed lines). All these estimates were obtained by using a regularization parameter $\mu_0 = 5 \times 10^{-2}$ (equation (19)). We can see, however, that even the estimate obtained with the full data set failed in retrieving the magnetization distribution of the simulated sample (red dots). This poorly estimated magnetization distribution yields predicted data (Figures 9b, 9e, 9h, and 9k) that do not recover the observed data (Figures 9a, 9d, 9g, and 9j). This coarse data fit is shown by the normalized histograms of the residuals between the predicted and observed data (Figures 9c, 9f, 9i, and 9l). The sample means $\mu$ would be very close to the simulated...
constant bias and the sample standard deviations $\sigma$ would be very close to that of the pseudo-random noise if the magnetic data were properly corrected from the misalignment problems.

These results exemplify the effects of not correcting the misalignment problems prior to the inversion. Besides, they show the advantage of inverting the full data set (data on the four planes). As we can see, these problems lead to an estimated magnetization distribution that do not retrieves the true one and also produce a coarse data fit.

### 3.1.3. Sensor-To-Sample Distance

Unlike the previous tests, the distance $h$ (Figure 3) between the surface of the sample and the observation points at each plane is not constant. On each plane $x = 0, 1, 2, 3$, the observation points are located on a set of 42 regularly spaced acquisition lines, each one with 102 regularly spaced points. In this test, the distance $h$ is contaminated with a Gaussian position noise having a common standard deviation of 100 $\mu$m and a different mean value equal to 420, 330, 400, and 230 $\mu$m on each plane $x = 0, 1, 2, 3$, respectively. The distance $h$, however, is constant for observation points located on the same acquisition line.

The noise-corrupted magnetic data were properly corrected from the misalignment problems by using the parameters shown in Table 2 and the equations (7) and (8). As in the previous tests, we applied our method by using an interpretation model with side lengths $L_x$, $L_y$, and $L_z$ equal to the true ones and with the same number of prisms ($P = 16$) as the synthetic sample along the $x$ axis. We have applied our method to invert the noise-corrupted magnetic data without, however, consider the variations in the sensor-to-sample distances $h$. Instead, we presumed a constant distance $h = 500$ $\mu$m between all the planes and the surface of the sample.

Figure 11 shows the four estimated magnetization distributions produced in this test. All these estimates were obtained by using a regularization parameter $\mu_0 = 1 \times 10^{-10}$ (equation (19)). This figure also shows that as in the previous tests, the estimate obtained by inverting the full data set (black dots and black dashed lines) best approximates the true magnetization distribution (red dots and red
Figure 9. Preprocessing errors test. Comparison between the true (red dots) and estimated (blue and black dots) magnetization (a) intensity, (b) inclination, and (c) declination. The estimated magnetization distribution obtained by inverting the magnetic data on the four planes $z = 0, 1, 2, 3$ is represented by the black dots and black dashed lines. The three sets of blue dots and dashed lines represent the estimated magnetization distribution obtained by inverting the magnetic data located on a single plane ($z = 0$), two planes ($z = 0, 1$), and three planes ($z = 0, 1, 2$) are plotted, respectively, with high, intermediate, and low transparency. The values are plotted along the $x$ axis, at the center of each prism forming the interpretation model. The black dashed lines in Figures 9b and 9c indicate the $\mu / C_{14}$ value.
Figure 10. Sensor-to-sample distance test. (a, d, g, j) Noise-corrupted magnetic data produced by the synthetic sample (not shown) on the observation planes \( z = 0, 1, 2, \) and 3, respectively. (b, e, h, k) Predicted data produced by the estimated magnetization distribution obtained by inversion on the observation planes \( z = 0, 1, 2, \) and 3, respectively. The estimated magnetization distribution was obtained by inverting the magnetic data located on the four observation planes. (c, f, i, l) Normalized histograms of the residuals between the predicted data shown in Figures 10b, 10e, 10h, and 10k and the noise-corrupted magnetic data shown in Figures 10a, 10d, 10g, and 10j. The normalization consists in subtracting from the residuals its sample mean \( \mu \) and dividing the result by its sample standard deviation \( \sigma \). The values are in \( \mu T \).
lines). However, even this estimate did not completely retrieve the magnetization distribution of the synthetic sample (red dots). Notice that the observed data (Figures 10a, 10d, 10g, and 10j) present striations aligned with the acquisition lines on each observation plane. These striations are produced by the presence of position noise in the observation points. Despite these striations, the (almost good) estimated magnetization distribution obtained by inverting the full data set (black dots and black dashed lines in Figure 11) yields predicted data (Figures 10b, 10e, 10h, and 10k) that recover the observed data (Figures 10a, 10d, 10g, and 10j) in an acceptable manner. This acceptable data fit is shown by the normalized histograms of the residuals between the predicted and observed data (Figures 10c, 10f, 10i, and 10l). The sample means $\mu$ are close to zero but the sample standard deviations $\sigma$ present some discrepancies with respect to that of the pseudo-random noise in magnetic data. According to these histograms, the poorest data fit occurs in the observation plane $x = 3$ (Figure 10l), where the sample standard deviation $\sigma \approx 47 \mu T$.

These results illustrate the effects of errors in the sensor-to-sample distance and also illustrate the advantage of inverting the full data set (data on the four planes). If compared with the results obtained in the previous test, the results obtained here suggest that our method is most robust against errors in the sensor-to-sample distance than against misalignment errors.

3.2. Marine Ferromanganese Crust Sample

Oda et al. [2011] proposed a new method for determining absolute ages and growth rates for ferromanganese crusts. They applied their method to analyze a ferromanganese crust sample obtained from a seamount in the Northwest Pacific Ocean and showed that magnetostratigraphy using state of the art SQUID microscopy is a promising chronological tool. The method proposed by Oda et al. [2011] uses SQUID microscopy on thin sections to identify the boundaries of fine magnetic stripes (smaller than 1 mm) with approximately reversed magnetization. These magnetic stripes are then correlated with a
Figure 12. Marine ferromanganese crust sample. (a, d, g, j) Noise-corrupted magnetic data produced by the synthetic sample (not shown) on the observation planes \( z = 0, 1, 2, \) and 3, respectively. (b, e, h, k) Predicted data produced by the estimated magnetization distribution obtained by inversion on the observation planes \( z = 0, 1, 2, \) and 3, respectively. The estimated magnetization distribution was obtained by inverting the magnetic data located on the four observation planes. The color scales are slightly saturated for improving the visualization. (c, f, i, l) Normalized histograms of the residuals between the predicted data shown in Figures 12b, 12e, 12h, and 12k and the noise-corrupted magnetic data shown in Figures 12a, 12d, 12g, and 12j. The normalization consists in subtracting from the residuals its sample mean \( \mu \) and dividing the result by its sample standard deviation \( \sigma \). The values are in nT.
standard magneto-stratigraphic time scale, providing a tool for estimating ages and growth rates for hydro-genetic ferromanganese crusts with unprecedented spatial resolution. Oda et al. [2011] also used SQUID magnetometer measurements to analyze a columnar block of their ferromanganese crust sample. They have cut and sliced the columnar block parallel to the growth lamination at 1.5 mm intervals using a 0.3 mm thick diamond-wire saw. After that, they estimated the magnetization of the slices by using a SQUID magnetometer. Note that to determine the magnetization along the sample, it was necessary to slice it in small parts. It would be useful instead to determine the bulk magnetization along the sample by using a nondestructive method. Here we present a feasibility study aiming at determining the magnetization distribution along a simulated sample. This feasibility study consists in applying our method and directly inverting simulated high-resolution measurements of the magnetic induction around the sample. The internal magnetization distribution of our simulated sample is based on the real magnetization distribution within the ferromanganese crust presented by Oda et al. [2011]. Our sample (not shown) is formed by $P_{24}$ juxtaposed prisms along the $x$ axis and has side lengths equal to $L_x = 36$ mm, $L_y = 5$ mm, and $L_z = 5$ mm along, respectively, the $x$, $y$, and $z$ axes. Figure 13 shows the magnetization of these prisms (red dots) along the $x$ axis of the sample. We have also simulated the presence of 20 magnetized grains that are randomly placed within the sample. These grains are represented by spheres having radius equal to 30 $\mu$m, constant magnetization intensity of 100 A/m and random magnetization direction. Notice that the presence of these magnetized grains violates the premise that the rock sample can be approximated by uniformly magnetized prisms. The magnetic data produced by our heterogeneous sample on each plane $z = 0, 1, 2, 3$ were calculated at a constant distance $h = 370$ $\mu$m, on a regular grid of $200 \times 100$ points along the $x$ and $y/z$ axes, respectively, and were also contaminated with a pseudo-random Gaussian noise having null mean and standard deviation equal to 0.1 mT. Figure 13 shows the magnetization of these prisms (red dots) along the $x$ axis of the sample.
standard deviation equal to 1.0 nT (Figures 12a, 12d, 12g, and 12j). The simulated magnetic sensor used for calculating these data has an square active area $A = (300 \mu m)^2$, which was discretized into a regular grid of $7 \times 7$ points, totaling $Q = 49$ (Figure 5).

Figure 13 shows the four estimated magnetization distributions produced in this test. All these estimates were obtained by using a regularization parameter $\mu_0 = 1 \times 10^{-10}$ (equation (19)). This figure also confirms that the estimate obtained by inverting the full data set (black dots and black dashed lines) is better than the remaining three (blue dots and blue dashed lines) because it completely retrieved the magnetization distribution of the simulated heterogeneous sample (red dots). This estimated magnetization distribution yields a predicted data (Figures 12b, 12e, 12h, and 12k) that is very close to the observed data (Figures 12a, 12d, 12g, and 12j), except at the regions where the simulated spherical grains are close to the sample surface. The normalized histograms of the residuals between the predicted and observed data (Figures 12c, 12f, 12i, and 12l) show sample means $\mu$ and sample standard deviations $\sigma$ very close to that of the pseudorandom noise contaminating the simulated magnetic data.

As in all the previous tests, these results confirm the advantage of inverting the full data set (data on the four planes). Moreover, they show the good performance of our method in retrieving the magnetization distribution within a complex and heterogeneous synthetic sample simulating a real ferromanganese crust.

4. Tentative Application to Real Data Produced by a Synthetic Sample

4.1. Sample Preparation

We have applied our method to invert experimentally measured data produced by a highly magnetized sample (not shown) that was manufactured in the laboratory of paleomagnetism of the Institute of Geophysics, Astronomy and Atmospheric Sciences of the University of São Paulo (IAG-USP), in Brazil. The sample is formed by four juxtaposed prisms with side lengths equal to $\approx 4$, $\approx 3$, and $\approx 3$ mm along, respectively, the $x$, $y$, and $z$ axis (with respect to the MCS shown in Figure 1). These prisms were manufactured by filling a rectangular acrylic mold with a magnetite solution and letting it hardens. The magnetite solution is formed by a mixture of a diamagnetic epoxy resin (0.3 g) and colloidal magnetite (0.02 g) obtained from the reaction of ferric sulfate and ferrous chloride. The magnetite has saturation magnetization $M_s$ ranging from 5 Am$^2$/kg and 8 Am$^2$/kg. The hardened prisms were then magnetized isothermally by an inducing field of $\approx 1$ T, according to the schematic representation shown in Figures 14a and 14b. Finally, the prisms were juxtaposed (Figure 14c), resulting in a sample with side lengths equal to $L_x \approx 16$ mm, $L_y \approx 3$ mm, and $L_z \approx 3$ mm. The inclination and declination values of the IRM within each prism are shown in Table 3. The resulting sample is referred to a MCS (Figure 1) with origin represented by the black dot and axes $x$ and $z$ represented by the thin arrows.
were measured at a constant distance combinations of parameters respectively (Figures 15a, 15d, 15g, and 15j).

LA60A, Zaber Technologies), allowing for the adjustment of the sensor to sample distance with micrometer effects. The sensor is mounted on another acrylic rod that can be raised or lowered using a linear actuator (T-
mounted to and increasing the distance from the sample to the micropositioners motors minimizing inductive

tes were made inside a small three-layer open-end magnetic shielded chamber (TLM S-0100, Bartington Instruments).
The experimental measurements were made inside a small three-layer open-end magnetic shielded chamber (TLM S-0100, Bartington Instruments).
The sensor has an active area size of 300 μm, which is 250 μm distant from the top of the encapsulation. The maximum spatial resolution achieved is about the size of the active area of the sensor. The magnetic field sensitivity measured is 350 nT/Hz$^{1/2}$ in the white noise region. The sensor is current biased and preamplified by a custom-made electronics at 1.0 kHz, and the output Hall voltage is detected by a lock-in amplifier. In our microscope, two independent linear microproportioners (T-LLS260, Zaber Technologies) were oriented perpendicularly and stacked, making up an x – y stage with a maximum travel range of 50 mm in each direction. An acrylic pedestal, 15 mm long, is fixed to the stage, serving two purposes, providing a flat surface for the sample to be mounted to and increasing the distance from the sample to the micropositioners motors minimizing inductive effects. The sensor is mounted on another acrylic rod that can be raised or lowered using a linear actuator (TLA60A, Zaber Technologies), allowing for the adjustment of the sensor to sample distance with micrometer accuracy. The scanning of the sample is made in a stop-and-go system, meaning that when the sample is moved to a scanning position, it stops for the magnetic field measurement and only then, it goes to the next position. In order to reduce environmental magnetic noise present at the laboratory, the experimental measurements were made inside a small three-layer open-end magnetic shielded chamber (TLM S-0100, Bartington Instruments).

We used for scanning a step of 200 μm in both x and y directions. On each plane $x = 0, 1, 2, 3$, the magnetic data were measured at a constant distance $h \approx 500 \mu m$, on a regular grid of $102 \times 42$ points along the x and y/z axes, respectively (Figures 15a, 15d, 15g, and 15j).

The observed magnetic data were corrected from the apparent misalignment problems by testing different combinations of parameters $l$, $\Delta x'$, and $\Delta y'$ (Figure 4b and equations (7) and (8)). The best parameters are very close to those ones shown in Table 2. After that, the magnetic data sets obtained on each plane were subtracted from their respective mean values.

### 4.3. Results and Comments

We inverted the observed magnetic data by using an interpretation model which is formed by $P = 16$ prisms and has side lengths equal to $L_x = 16$ mm, $L_y = 3$ mm, and $L_z = 3$ mm along, respectively, the x, y, and z axes. Note that each prism forming the synthetic sample is approximated by four juxtaposed prisms of the interpretation model. We considered a square magnetic sensor with active area $A = (300 \mu m)^2$, which was discretized into a regular grid of 7 × 7 points, totaling $Q = 49$ (Figure 5). Figure 16 shows the four estimated magnetization distribution obtained with our method. All these estimates were obtained by using a regularization parameter $\mu_0 = 1 \times 10^{-10}$ (equation (19)). The black dots and black dashed lines represent the estimate obtained by inverting the magnetic data on the four observation planes. This

<table>
<thead>
<tr>
<th>Index</th>
<th>$l$ (°)</th>
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<tr>
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*IRM inclination ($l$) and declination ($\theta$) of the prisms forming the synthetic sample that was manufactured in laboratory. Each prism is indicated by an index, according to the Figure 14c.
Figure 15. Application to real data. (a, d, g, j) Observed magnetic data produced by the synthetic sample (not shown) on the observation planes \( \alpha = 0, 1, 2, \) and 3, respectively. (b, e, h, k) Predicted data produced by the estimated magnetization distribution obtained by inversion on the observation planes \( \alpha = 0, 1, 2, \) and 3, respectively. The estimated magnetization distribution was obtained by inverting the magnetic data located on the four observation planes. (c, f, i, l) Normalized histograms of the residuals between the predicted data shown in Figures 15b, 15e, 15h, and 15k and the observed magnetic data shown in Figures 15a, 15d, 15g, and 15j. The normalization consists in subtracting from the residuals its sample mean \( \mu \) and dividing the result by its sample standard deviation \( \sigma \). The values are in \( \mu T \).
estimated magnetization distribution yields a predicted data (Figures 15b, 15e, 15h, and 15k) that recover the main features in the observed data (Figures 15a, 15d, 15g, and 15j). This coarse data fit is shown by the normalized histograms of the residuals between the predicted and observed data (Figures 15c, 15f, 15i, and 15l). These histograms show nonnull sample means $\mu$ and different sample standard deviations $\sigma$.

Based on the results obtained with synthetic data, this poor data fit may be related to the misalignment problems. Probably, the parameters $h$, $D_x^0$, and $D_y^0$ (Figure 4b and equations (7) and (8)), which were estimated by trial-and-error, do not lead to a suitable correction of the misalignment problems. These problems, however, were possibly worsened by the apparent imprecision in the sensor positioning (or position noise) during the scanning stage. Notice that as in the numerical test showing the effects of position noise, the observed data (Figures 15c, 15f, 15i, and 15l) present striations along the acquisition lines on each observation plane. According to Lee et al. (2004), the presence of position noise has an important implication for the design of scanning magnetic microscopy. We have also verified with synthetic data that errors in the sensor-to-sample distance could negatively impact the results obtained by our method.

These results cannot nonetheless be completely rejected owing to the estimated magnetization distribution along the synthetic sample shows two important characteristics.

The first one concerns the estimated magnetization intensities (Figure 16a). As we can see, the estimated intensities exhibit an interesting pattern formed by a well-defined cycle of a low value that is followed by three higher values. This cycle repeats four times, coinciding with the number of prisms forming our synthetic sample. For convenience, the estimated values

![Figure 16. Application to real data. Estimated magnetization (a) intensity, (b) inclination, and (c) declination. The estimated magnetization distribution obtained by inverting the magnetic data on the four planes $z = 0, 1, 2, 3$ is represented by the black dots and black continuous lines. The three sets of blue dots and dashed lines represent the estimated magnetization distribution obtained by inverting the magnetic data located on a single plane ($z = 0$), two planes ($z = 0, 1$), and three planes ($z = 0, 1, 2$) are plotted, respectively, with high, intermediate, and low transparency. The values are plotted along the $x$ axis, at the center of each prism forming the interpretation model. The continuous (vertical) black lines divide the estimated values representing each prism forming the sample. The numbers indicate the index of each prism (Figure 14c and Table 3). The black dashed (horizontal) lines in Figure 16b indicate the values $0°$, $45°$, and $90°$. The black dashed (horizontal) lines in Figure 16c indicate the values $0°$ and $180°$. The estimated values that are represented by black triangles are considered spurious due to the magnetite precipitation.](image-url)
are grouped to represent the magnetization within each prism. The low values observed in the estimated magnetization intensities (Figure 16a), as well as the corresponding inclination and declination values (Figures 16b and 16c) are represented by black triangles. This remarkable pattern is consistent with the magnetite precipitation during the manufacturing of our sample. This precipitation would decrease the magnetite concentration at the top of each sample prism, increasing the concentration toward their bottom. Figures 14b and 14c represent the magnetite precipitation as a gradient fill. The resultant high-concentration zones at the bottom of the prisms forming the sample are represented in dark gray while the resultant low-concentration zones at the top of the prisms are represented in white. The magnetic induction produced by these low-concentration zones there must be superimposed by those ones produced by the high-concentration zones, which makes the correct estimation of the magnetization direction in these low-concentration zones very difficult. Due to this possible precipitation effect, we did not take into consideration the estimated values that coincide with the low-concentration zones (black triangles in Figure 16). Consequently, we based our interpretation on the estimated values that are not located at the low-concentration zones.

The second interesting characteristic of the estimated magnetization distribution along the synthetic sample is about the estimated inclination/declination (Figures 16b and 16c). Note that the estimated inclinations (blue dots in Figure 16b) within the four prisms forming the sample are close to the approximated values shown in Table 3. On the other hand, the estimated declinations (blue dots in Figure 16c) are close to the values shown in Table 3 at the prisms 0, 1, and 2, but are very different at the prism 3. This discrepancy nonetheless is not a problem because, at prism 3, the estimated inclination is close to 90° (Figure 16b) and, in this case, the estimated declinations are not important.

These results show that even inverting a magnetic data set that were measured by a prototype magnetometer and were, at least apparently, highly contaminated with position noise, our method estimated a magnetization distribution that is consistent with the IRM orientation within our manufactured sample.

5. Conclusion

We have presented a new method for inverting scanning magnetic microscopy data with the purpose of estimating the internal magnetization distribution of a rectangular rock sample whose internal magnetization distribution varies along one of its axes. Our method presents a new approach for automatically including the averaging effect produced by the finite size of the magnetic sensor. Additionally, our method takes advantage of the geometry of a rectangular sample to propose a new scanning design aiming at providing the magnetic data not only on a single plane over the sample, but on four mutually orthogonal planes located around the sample. The use of magnetic data located on the four planes instead of a single one introduces independent information into the linear system to be solved and consequently leads to a better conditioned inverse problem. The advantage of inverting magnetic data located on four planes around the sample is confirmed by all tests with synthetic data.

Results with synthetic data produced by a simulated highly magnetized sample show not only the good performance of our method in retrieving the magnetization distribution within a rock sample in an ideal case, but also how the results obtained by our method are negatively impacted by imprecisions in the sensor positioning (or position noise). Results with synthetic data produced by a simulated sample that resembles a marine ferromanganese crust having a complex and heterogeneous internal magnetization distribution suggest that our method could be applied to invert magnetic data produced by real geological samples. We have also applied our method to invert experimental magnetic data for estimating the internal magnetization distribution of a synthetic sample that was manufactured in laboratory. Despite the misalignment problems occurred during the scanning stage and also the apparent high position noise in the experimental magnetic data, our method estimated a meaningful magnetization distribution that is consistent with that in the manufactured sample.

These results show that our method could be an interesting complement to traditional paleomagnetic techniques aiming at providing a nondestructive diagnostic of geological samples. Nevertheless, further tests using either manufactured and geological samples need to be carried out with the purpose of validating our approach, better evaluating the effects of position noise and misalignment problems during the scanning stage and also developing automatic preprocessing techniques. A further study need to be carried out
in order to evaluate the influence of, for example, the scanning geometry and the sensitivity of the magnetic sensor on the performance of our method.

By presuming that the internal magnetization of the rectangular sample varies along one of its axes, we implicitly restrict the application of our method to samples that are perpendicicular to the growth lamination within rocks. It would be interesting to generalize our method in order to remove this restriction and allow its application to estimate complex magnetization distributions within rock samples.

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